ARPN Journal of Engineering and Applied Sciences

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STOCHASTIC MODELING OF MONTHLY RAINFALL AT KOTA REGION

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ABSTRACT

A study on stochastic modeling for rainfall was undertaken using 35 years (1987–2003) data of Kota region. The performed statistical tests indicated that the series of the monthly rainfall data was trend free. The periodic component of monthly rainfall could be represented by third harmonic expression. The stochastic components of the monthly rainfall followed first order Markov model. Validation of generated monthly rainfall series was done by comparison of generated and measured series. The correlation coefficient between generated and measured rainfall series was found to be 0.9961. The correlation was tested by t-test and found to be highly significant at 1 per cent level. The regression equation is very near to 1:1 line. Therefore, developed model could be used for future prediction of monthly rainfall.

Keywords: stochastic, modeling, rainfall, kota.

INTRODUCTION

Rainfall is an important weather parameter for estimation of crop water requirements. Frequently, it is required to estimate rainfall of places where measured rainfall data are not available. Rainstorms vary greatly in space and time. This type of series could be estimated by the sum of periodic series and stochastic series. Periodic component takes into account the portion, which repeats after certain duration. The stochastic component constituted by various random effects, can not be estimated exactly.

MATERIALS AND METHODS

Study area

The study area comes under the sub-humid region of the agro-climatic zone V of the state of Rajasthan, and is situated at 25°11' N latitude, 75°51' E longitude and at an altitude of 256.9 m above mean sea level. The annual rainfall in this region is 761.1 mm and more than 89% of this amount is received during the monsoon season alone, due to the influence of the southwest monsoon.

Collection of rainfall data

The rainfall data of Kota region was collected from Central Soil and Water Conservation Research and Training Institute, Research Centre, Kota. Rainfall data for a period of 35 years (1970-2004) was used in the study.

The mathematical procedure adopted for formulation of a predictive model has been discussed as follows: The principal aim of the analysis was to obtain a reasonable model for estimating the generation process and its parameters by decomposing the original data series into its various components. Generally a time series can be decomposed into a deterministic component, which could be formulated in manner that allowed exact prediction of its value, and a stochastic component, which is always present in the data and can not strictly be

accounted for as it is made by random effects. The time series, $X_{(t)}$, was represented by a decomposition model of the additive type, as follows:

$$X_{(t)} = T_{(t)} + P_{(t)} + S_{(t)}$$
 ...(1)

where

 $T_{(t)}$ = trend component, t = 1,2N

 $P_{(t)}$ = periodic component

 $S_{(t)}$ = stochastic component, including dependent and independent parts.

To obtain the representative stochastic model of time series, identification and detection of each component of Equation (1) was necessary. A systematic identification and reduction of each component of $X_{(t)}$ was done, procedures of which are described below:

Trend component

The trend component describes the long smooth movement of the variable lasting over the span of observations, ignoring the short term fluctuations. The basic idea here was to study only $T_{(t)}$ while eliminating the effects of other components. It leads to use the total seasonal data, Z_t , for identification of $T_{(t)}$ so that other components were suppressed. For detecting the trend, a hypothesis of no trend was made. Turning point test and Kendall's rank correlation tests as suggested by Kottegoda (1980) were performed. If the calculated value of Z is within its table value, then, it can be concluded that the trend is not present in the series.

Periodic component

The periodic component concerns an oscillating movement which is repetitive over a fixed interval of time (Kottegoda 1980). The existence of $P_{(t)}$ was identified by the correlogram, a plot of autocorrelation coefficients, r_1 , versus lag 1. The oscillating shape of the correlogram verifies the presence of $P_{(t)}$, with the seasonal period $P_{(t)}$, at the multiples of which peak of estimation can be made by a Fourier Analysis followed by the tests for significant harmonics. The correlogram of the time series

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clearly showed the presence of the periodic variations indicating its detection. The periodic component $X_{(t)}$ was expressed in Fourier form as follows:

$$X_{(t)} = A_o + \sum_{K=1}^{\infty} \left[A_K \cos\left(\frac{2\pi Kt}{p}\right) + B_K \sin\left(\frac{2\pi Kt}{p}\right) \right] \dots (2)$$
where,
$$A_o = \frac{1}{N} \sum_{t=1}^{N} x_{(t)} \dots (3)$$

$$A_K = \frac{2}{N} \sum_{t=1}^{N} x_{(t)} \cos\left(\frac{2\pi Kt}{p}\right) \dots (4)$$
and
$$B_K = \frac{2}{N} \sum_{t=1}^{N} x_{(t)} \sin\left(\frac{2\pi Kt}{p}\right) \dots (5)$$

K = number of significant harmonics, p = baseperiod, N = number of observation points and A_K and B_K = Fourier coefficients.

These coefficients were obtained by a least square fit of the data to the Kth harmonic components, then a least squares approximation could be given by the finite series.

$$P_{(t)} = A_o + \sum_{K=1}^{M} \left[A_K \cos\left(\frac{2\pi Kt}{p}\right) + B_K \sin\left(\frac{2\pi Kt}{p}\right) \right] \dots (6)$$

Where, M is the number of significant harmonics (maximum, P/2). In Equation (6) if $M \rightarrow \infty$, $P_{(t)} \rightarrow X_{(t)}$, then X_(t) could be represented satisfactorily by Equation (6) only. However it may not be practical or desirable to allow the condition, $M \rightarrow \infty$. Thus the appropriate approach would be the selection of a value of M which contains only those harmonics which are significantly contributing towards $X_{(t)}$. With this as the objective two tests namely test of analysis of variance and Fourier decomposition of mean square were conducted.

Stochastic component

The stochastic component was constituted by various random effects, which could not be estimated exactly. In the case of rainfall time series various climatic parameters response to value of component without changing the cyclicity itself and thus add randomness to the time series. A stochastic model in the form of Autoregressive model, AR, was used for the presentation in the time series. In this model, the current value of the process was expressed as a finite, linear aggregate of values of the process and a variate that was completely random. This model was applied to the $S_{(t)}$ which was treated as a random variable i.e. deterministic components were removed and the residual was stationary in nature. Mathematically, an autoregressive model of order p, AR (p) can be written as:

$$S_{(t)} = \sum_{K=1}^{P} \phi_{P,K} S_{(t-K)} + a_{(t)}$$

$$= \phi_{p,I} S_{(t-1)} + \phi_{p,2} S_{(t-2)} + \dots + \phi_{p,p} S_{(t-P)} + a_{(t)}$$
(7)

Where, $\phi_{P,K}$ = autoregressive model parameters,

 $a_{(t)}$ = independent random number

The fitting procedures of the AR(p) model to the meteorological parameters series involved selection of order (p) of the model.

Estimation of autoregressive parameters

The parameter estimation deals with the estimation of the autoregressive parameters of Equation (7). These parameters could be expressed in terms of serial correlation coefficient, as Yule-Walker equations (Bhakar, 2000). The general recursive formulae for estimating these parameters $(\phi_{p,k})$, where suffix p and k indicate the order and the number of parameter of the order in AR (p) model, respectively could be written as follows:

$$\phi_{p,p} = \left[\frac{r_p - \sum_{k=1}^{p-1} (\phi_{p-1,k}) (\cdot k_{p-k}^6)}{1 - \sum_{k=1}^{p-1} (\phi_{p-1,k}) (r_k)} \right] ...(8)$$

$$\phi_{p,k} = \phi_{p-1,k} - \phi_{p,p}. \phi_{p-1,p-k}; k = 1,2,3....p-1$$
 ...(9)

In Equation (8) r_k is the autocorrelation coefficient, Auto Correlation Coefficient of the series for K and was computed, for any series $Y_{(t)}$ at any lag, l, as follows:

$$r_{1} = \sum_{t=1}^{N-1} [Y_{(t)} - Y_{(t)}][Y_{(t+1)} - Y_{(t)}] / \sum_{t=1}^{N} [Y_{(t)} - Y_{(t)}]^{2} = C_{1} / C_{0}$$
.....(10)

Where, $Y_{(t)} = \text{mean of the series, } Y_{(t)}$ $N = \text{total number of discrete values of } X_{(t)}$ $C_l = \text{autocovariance function at lag } l, l = 0, 1, ..p$

After estimating the AR parameters $\phi_{p,k}$, $S_{(t)}$ was calculated by using Equation (7). The sum of the periodic and stochastic component forms the generated value of the observed data. The difference was termed as residuals which were tested to check the adequacy of the formulated model.

Diagnostic checking of the model

Diagnostic checking concerns the verification for the adequacy of the fitted model. The examination of the autocorrelation structure of the residuals provides a powerful way of diagnostic checking. The residuals were examined for any lack of randomness. If the residuals were not random or were autocorrelated, the model has to be modified until the residuals become uncorrelated.

RESULTS AND DISCUSSION

VOL.1, NO.3, OCTOBER 2006

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For testing the statistical characteristics of monthly rainfall series, 35 years data of monthly rainfall was taken. Average of 35 was taken for every week to get mean monthly rainfall series. The statistical characteristics of the mean monthly rainfall series were estimated. Mean monthly rainfall values vary from 2.71 mm in the month of March to 254.72 mm in the month of July. Mean values was found to be 60.51 mm. Mean values was found to be 60.51 mm. There is large variability among the monthly values of rainfall of different years. The coefficient of variation varies from 0.389 in the month of August to 3.101 in the month of November. Variance is maximum during monsoon season. It indicates atmospheric instability during monsoon. Minimum value of standard deviation is 6.025mm during March. It indicates better weather stability during month of March. The variation may be attributed towards the natural changes in yearly climate. Annual rainfall of Banswara indicates minimum 201.1mm during 2003 AD and Maximum 1528.3mm in the year of 1996 AD with mean rainfall 761.1mm.

Serial correlation coefficient

The lag one serial correlation coefficient of observed series was calculated by Equation (10) and was found to be 0.457. The respective confidence limits were estimated as -0.104 to 0.099 (Kottegoda, 1980). The value of lag one serial correlation coefficient lies outside the range of confidence limits and is significantly different from zero. This again confirms that rainfall process is mutually dependent. From the analysis of coefficient of variation and serial correlation, it is confirmed that rainfall process is time variant and not an independent one. Thus the rainfall time series could be modeled on stochastic theory. The mutual dependence of the observed rainfall series was also confirmed by the correlogram (Figure-1).

Trend component

For identification of trend components, annual rainfall series was used. The annual rainfall series was obtained by transforming the 35 years annual series. For detection of trend the hypothesis of no trend was made and value of test statistics (Z) was calculated by Turning Point Test and Kendall's Rank Correlation Test. The calculated values of test statistics (Z) are -0.28 and -1.77 for Turning Point Test and Kendall's Correlation Test respectively. The estimated value of test statistics (Z) was within the 1 percent level of significance. Hence the hypothesis of no trend was accepted. From the above analysis it is confirmed that the trend component in rainfall time series is absent and the observed series may be treated as trend free series.

Periodic component

$$P_{t} = 60.51 - 71.837 \cos\left(\frac{2\pi t}{p}\right) - 69.882 \sin\left(\frac{2\pi t}{p}\right) + 6.389 \cos\left(\frac{4\pi t}{p}\right) + 68.265 \sin\left(\frac{4\pi t}{p}\right)$$

To confirm the presence of periodic component in monthly rainfall series a correlogram was drawn. A correlogram is a graphical representation of serial correlation coefficients (r_l) as function of lag l in which the values of r_l are plotted against respective value of l (Figure-1). The resulting oscillating shape of the correlogram confirms the presence of periodic component in the monthly rainfall. Further, the correlogram has peaks at legs equal to 12 and at other multiples of it. The time span of periodicity was taken as 12 for use in harmonic analysis of periodic component.

Determination of significant harmonics

For representing the periodic component of the rainfall series the numbers of significant harmonics were determined by analyzing by cumulative periodogram. Only first three harmonics are highly significant. Other harmonics were not significant and therefore could be ignored.

Parameters of periodic component

Using Equations (3), (4) and (5) the Fourier coefficients A_k and B_k were estimated. Study of which reveals that the first three harmonics explain more than 66 percent of variance.

Cumulative periodogram test

In this test a graphical procedure has been employed as criteria for obtaining the significant harmonics to be fitted in a periodic component. A graph was drawn between Pi and number of harmonics, called the cumulative periodogram. The fast increase in Pi has been considered as a significant harmonics and the rest of harmonics were rejected.

Estimates of the mean square deviation and the cumulative periodogram Pi for monthly rainfall series were made and the plot of Pi against i is shown in Figure-2. From the cumulative periodogram in the monthly rainfall series, it can be observed that the first three harmonics appeared to be the periodic part of the fast increase and after that periodogram remains almost constant which may be treated as non-significant.

The two criteria used to identify the number of significant harmonics to be used in modeling periodic component were found to be consistent, so first three harmonics are treated as significant contributing towards periodicity and remaining are considered as white noise.

For the first three harmonics the values of Fourier coefficients (A₁, A₂, A₃, B₁, B₁, B₃) were found to be -71.837, 6.389, 21.757 and -69.882, 68.265, -32.761, respectively. With these coefficients and using Equations (6) the periodic component (P_t) resulting from periodic deterministic process may be mathematically expressed

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$$+21.757\cos\left(\frac{6\pi t}{p}\right)-32.761\sin\left(\frac{6\pi t}{p}\right)$$

The deterministic cycle component (P_t) was computed by using Equation (11). After estimating the periodic component it was removed from, historical time series by subtracting the periodic component from historical time series. This process obtained a new stationery series, S_t , resulting from stochastic non-deterministic process.

Stochastic component

The presence of stochastic component was already confirmed by plotting the correlogram (Figure-1) of observed series and analysis of serial correlation coefficient (SCC) and coefficient of variance (CV). The periodic component was removed from the historical series. The rest of the data were analyzed to obtain non-deterministic stochastic component by fitting the autoregressive or Markov process of stochastic modeling.

Selection of model order

Residual variance method was used to determine the order of the model which may significantly represent the non-deterministic stationary stochastic component. Residual variance at different lags was computed. The minimum residual, variance was obtained for order one. ...(11)

The values of residual variance after 1st order showed no definite trend.

Using Equation (7) and the estimated autoregression coefficients the stochastic component of the monthly rainfall time series may be expressed as:

$$S_t = 0.457 S_{t-1} + a_t$$
 ...(12)

Where S_t and S_{t-1} are stochastic component at time t and (t-1).

Residual series of stochastic component

The residual series (a_t) which is random independent part of stochastic component was obtained after removing the periodic and dependent stochastic parts from the historical series.

The statistical analysis of the residual series confirms its normal distribution with mean which is almost equal to zero (mean = -0.01 and SD = 35.111). The values of statistical measures are presented in Table-1. The mean, SD of the historical and generated series are almost same which shows closeness between historical and generated data.

Table-1. Statistical parameters of the observed, generated and residual series of monthly rainfall.

Parameters	Historical series	Generated series	Residual series
Mean, mm	62.34	62.36	-0.01
SD, mm	111.083	95.066	35.111
Variance	12339.454	9037.480	1232.785

Model structure

Since the observed monthly rainfall series was found to be a trend free series that developed model describes the periodic-stochastic behaviour of the series.

The developed model is a superimposition of third harmonic deterministic process and second order autoregressive model. The mathematical structure of the additive model can now be represented as follows:

$$Rain = 60.51 - 71.837 \cos\left(\frac{2\pi t}{p}\right) - 69.882 \sin\left(\frac{2\pi t}{p}\right) + 6.389 \cos\left(\frac{4\pi t}{p}\right) + 68.265 \sin\left(\frac{4\pi t}{p}\right) + 21.757 \cos\left(\frac{6\pi t}{p}\right) - 32.761 \sin\left(\frac{6\pi t}{p}\right) + 0.457 S_{t-1} + a_t \qquad ...(13)$$

The first seven terms in the formulated model represented by Equation (13) constitute the deterministic part of the monthly rainfall time series. The eighth term represents the dependent stochastic component of the model where the current value of S_t depends on the observed preceding one value. The last term is the random independent part of the stochastic component. Using the developed model the average monthly rainfall series was generated for all the values.

Diagnostic checking of rainfall model

The residuals obtained after fitting the formulated model was subjected to various analysis to

test their adequacy for representing the time dependent structure of the monthly rainfall.

Sum of squares analysis

The sum of squares of residuals series were compared with sum of squares of deviations of observed values from their mean. The value of coefficient of determination (R²) was found to be 0.9001. Thus, this leads to the conclusion that the developed model has a fair goodness of fit to generate the monthly rainfall series.

VOL.1, NO.3, OCTOBER 2006 ISSN 1819-6608

ARPN Journal of Engineering and Applied Sciences

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Serial correlation analysis

The serial correlation coefficients (SCC) for lags l (l=1, 2, 3....104) were computed with the help of Equation (10). The values of SCC against respective lags were then plotted to obtain a correlogram. The resulting correlogram is shown in Figure 3 with confidence limit at 1 percent level. The correlogram is almost completely contained within the confidence limits at 1 per cent level. Hence it may be treated to be non-significant. This again confirms that the residual series may be treated as random series.

The residual series has a mean value of -0.01 and the variance of 1232.785. This leads to the conclusion that the residuals are independent normally distributed. Further, it also confirms the randomness of the residuals.

Validation of Stochastic Model of Monthly Rainfall

Validation of generated monthly rainfall series by developed stochastic model (Equation 13) was done by comparison of generated monthly rainfall series and measured monthly rainfall series. Validation of generated 35 year mean monthly rainfall series was made with 35 year mean measured rainfall series (Figure-4). The relationship is shown in Figure-5. The correlation coefficient between generated mean monthly rainfall series and measured mean monthly rainfall was found to be 0.9961. The correlation was tested by t test and found to be highly significant at 1 percent level. The standard error (8.82mm) is quite low. The mean of the monthly generated rainfall was found to be 62.4mm. Mean of the measured monthly rainfall series was found to be 62.3. The regression equation is very near to 1:1 line.

Therefore, Equation (13) can be used for future prediction of monthly rainfall.

Validation of generated 2003 and 2004 monthly rainfall series was made with measured monthly rainfall series (Figure-6). The relationship between measured monthly rainfall and estimated monthly rainfall is shown in Figure-7 (for year 2003). The correlation coefficient values between generated monthly rainfall series and measured monthly rainfall were found to be 0.7344 and 0.9481 for the years 2003 and 2004 respectively. The correlation was tested by t test and found to be highly significant at 1 percent level. The regression equations are near to 1:1. Therefore, Equation (13) could be used for future prediction of monthly rainfall values under climatic conditions of Kota.

ACKNOWLEDGEMENT

This paper is part of the research work carried out under the project entitled "Stochastic modeling of agro-meteorological phenomena for the farm water management planning" funded by the ICAR, Govt. of India.

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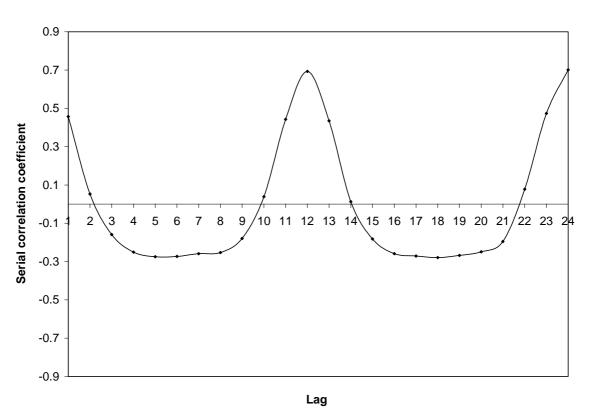


Figure-1. Correlogram of annual rainfall at Kota.

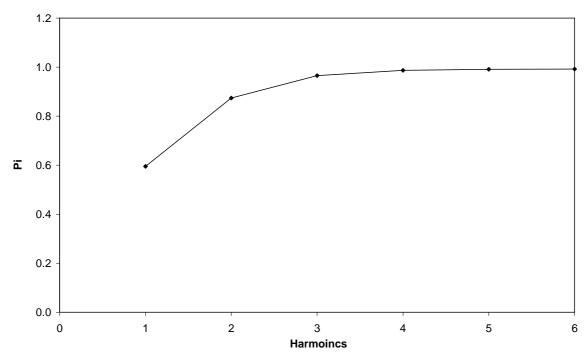


Figure-2. Cumulative periodogram of monthly rainfall.



ISSN 1819-6608

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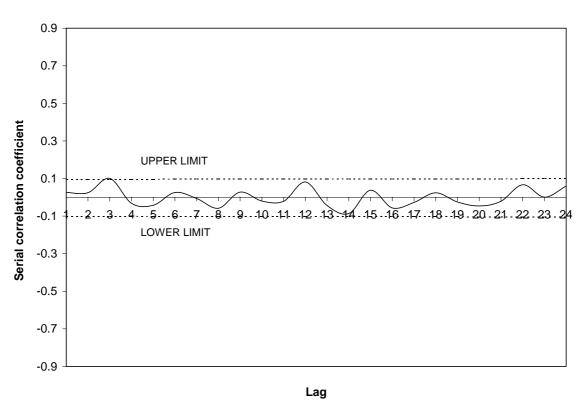


Figure-3. Correlogram up to lag 24 for residual series of monthly rainfall for 35 years (1970-2004) at Kota.

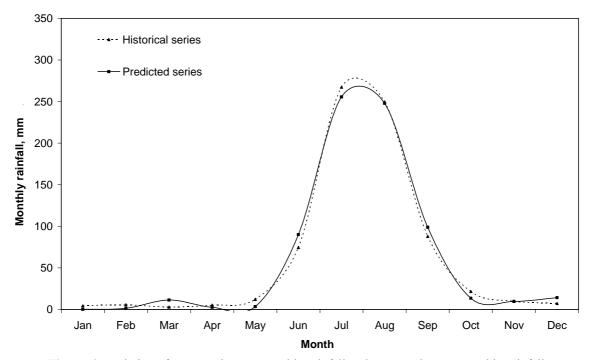


Figure-4. Variation of generated mean monthly rainfall and measured mean monthly rainfall for 33 years (1970-2002) at Kota.

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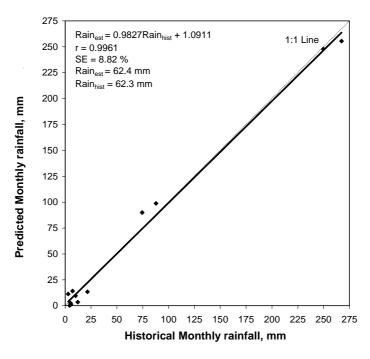


Figure-5. Relationship between generated mean monthly rainfall (Rainest) and measured mean monthly rainfall (Rainhist) for 33 years (1970-2002) at Kota.

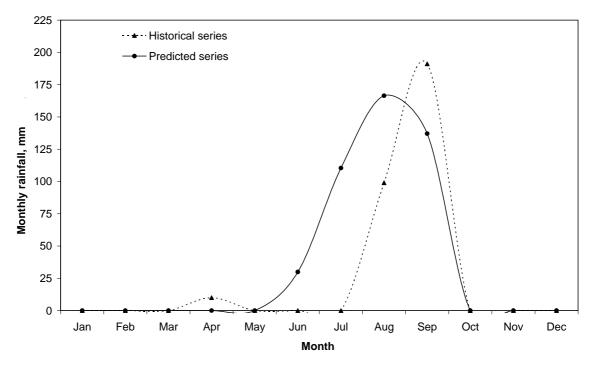


Figure-6. Variation of generated mean monthly rainfall and measured mean monthly rainfall for year 2003 at Kota.

VOL.1, NO.3, OCTOBER 2006 ISSN 1819-6608



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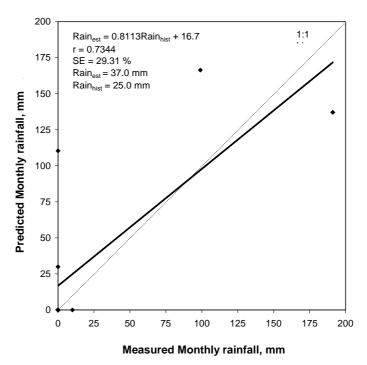


Figure-7. Relationship between generated mean monthly rainfall (Rainest) and measured mean monthly rainfall (Rainhist) for 2003 at Kota.